Inference in Bayesian networks

AIMA2E CHAPTER 14.4-5

Outline

- ♦ Exact inference by enumeration
- \Diamond Exact inference by variable elimination
- ♦ Approximate inference by stochastic simulation
- \Diamond Approximate inference by Markov chain Monte Carlo

Inference tasks

Simple queries: compute posterior marginal $P(X_i|\mathbf{E} = \mathbf{e})$ e.g., P(NoGas|Gauge = empty, Lights = on, Starts = false)

Conjunctive queries: $P(X_i, X_j | \mathbf{E} = \mathbf{e}) = P(X_i | \mathbf{E} = \mathbf{e})P(X_j | X_i, \mathbf{E} = \mathbf{e})$

Optimal decisions: decision networks include utility information; probabilistic inference required for P(outcome|action, evidence)

Value of information: which evidence to seek next?

Sensitivity analysis: which probability values are most critical?

Explanation: why do I need a new starter motor?

Inference by enumeration

Slightly intelligent way to sum out variables from the joint without actually constructing its explicit representation

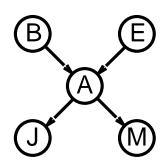
Simple query on the burglary network:

$$\mathbf{P}(B|j,m)$$

$$= \mathbf{P}(B,j,m)/P(j,m)$$

$$= \alpha \mathbf{P}(B,j,m)$$

$$= \alpha \Sigma_e \Sigma_a \mathbf{P}(B,e,a,j,m)$$



Rewrite full joint entries using product of CPT entries:

$$\begin{aligned} &\mathbf{P}(B|j,m) \\ &= \alpha \Sigma_e \Sigma_a \mathbf{P}(B) P(e) \mathbf{P}(a|B,e) P(j|a) P(m|a) \\ &= \alpha \mathbf{P}(B) \Sigma_e P(e) \Sigma_a \mathbf{P}(a|B,e) P(j|a) P(m|a) \end{aligned}$$

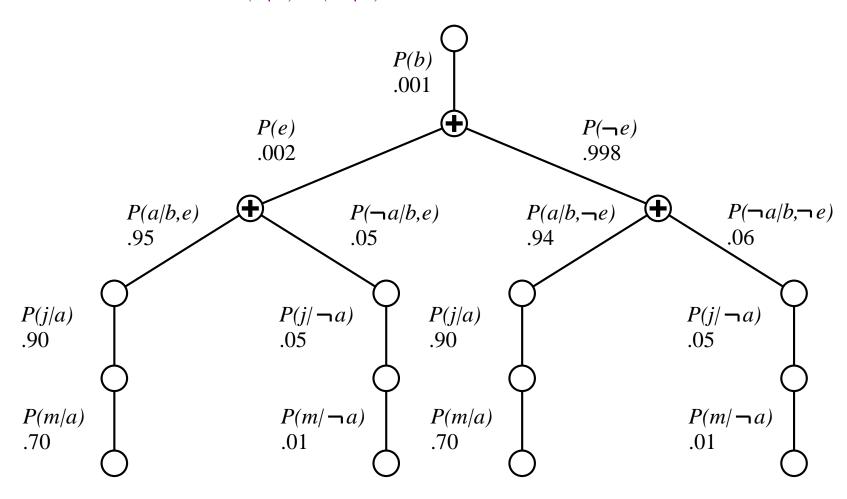
Recursive depth-first enumeration: O(n) space, $O(d^n)$ time

Enumeration algorithm

```
function ENUMERATION-ASK(X, e, bn) returns a distribution over X
   inputs: X, the query variable
              e. observed values for variables E
              bn, a Bayesian network with variables \{X\} \cup \mathbf{E} \cup \mathbf{Y}
   \mathbf{Q}(X) \leftarrow a distribution over X, initially empty
   for each value x_i of X do
         extend e with value x_i for X
         \mathbf{Q}(x_i) \leftarrow \text{ENUMERATE-ALL}(\text{VARS}[bn], \mathbf{e})
   return Normalize(\mathbf{Q}(X))
function ENUMERATE-ALL(vars, e) returns a real number
   if EMPTY?(vars) then return 1.0
   Y \leftarrow \text{First}(vars)
   if Y has value y in e
        then return P(y \mid Pa(Y)) \times \text{ENUMERATE-ALL(REST(vars), e)}
         else return \Sigma_y P(y \mid Pa(Y)) \times \text{Enumerate-All(Rest(vars), } \mathbf{e}_y)
              where \mathbf{e}_{y} is \mathbf{e} extended with Y = y
```

Evaluation tree

Enumeration is inefficient: repeated computation e.g., computes P(j|a)P(m|a) for each value of e



Inference by variable elimination

Variable elimination: carry out summations right-to-left, storing intermediate results (factors) to avoid recomputation

$$\begin{aligned} \mathbf{P}(B|j,m) &= \alpha \underbrace{\mathbf{P}(B)}_{B} \underbrace{\sum_{e} P(e)}_{E} \underbrace{\sum_{a} \mathbf{P}(a|B,e)}_{A} \underbrace{P(j|a)}_{J} \underbrace{P(m|a)}_{M} \\ &= \alpha \mathbf{P}(B) \underbrace{\sum_{e} P(e)}_{E} \underbrace{\sum_{a} \mathbf{P}(a|B,e)}_{A} P(j|a) f_{M}(a) \\ &= \alpha \mathbf{P}(B) \underbrace{\sum_{e} P(e)}_{a} \mathbf{P}(a|B,e) f_{J}(a) f_{M}(a) \\ &= \alpha \mathbf{P}(B) \underbrace{\sum_{e} P(e)}_{E} \underbrace{\sum_{a} f_{A}(a,b,e)}_{J} f_{J}(a) f_{M}(a) \\ &= \alpha \mathbf{P}(B) \underbrace{\sum_{e} P(e)}_{\bar{A}JM}(b,e) \text{ (sum out } A\text{)} \\ &= \alpha \mathbf{P}(B) f_{\bar{E}\bar{A}JM}(b) \text{ (sum out } E\text{)} \\ &= \alpha f_{B}(b) \times f_{\bar{E}\bar{A}JM}(b) \end{aligned}$$

Variable elimination: Basic operations

Summing out a variable from a product of factors: move any constant factors outside the summation add up submatrices in pointwise product of remaining factors

$$\sum_{x} f_1 \times \cdots \times f_k = f_1 \times \cdots \times f_i \sum_{x} f_{i+1} \times \cdots \times f_k = f_1 \times \cdots \times f_i \times f_{\bar{X}}$$

assuming f_1, \ldots, f_i do not depend on X

Pointwise product of factors f_1 and f_2 :

$$\begin{split} f_1(x_1,\dots,x_j,y_1,\dots,y_k) \times f_2(y_1,\dots,y_k,z_1,\dots,z_l) \\ &= f(x_1,\dots,x_j,y_1,\dots,y_k,z_1,\dots,z_l) \\ \text{E.g., } f_1(a,b) \times f_2(b,c) &= f(a,b,c) \end{split}$$

Variable elimination algorithm

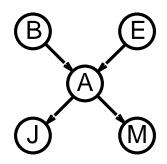
```
function ELIMINATION-ASK(X, e, bn) returns a distribution over X inputs: X, the query variable e, evidence specified as an event bn, a belief network specifying joint distribution \mathbf{P}(X_1,\ldots,X_n) factors \leftarrow []; \ vars \leftarrow \text{Reverse}(\text{Vars}[bn]) for each var in vars do factors \leftarrow [\text{Make-Factor}(var,\mathbf{e})|factors] if var is a hidden variable then factors \leftarrow \text{Sum-Out}(var,factors) return Normalize(Pointwise-Product(factors))
```

Irrelevant variables

Consider the query P(JohnCalls|Burglary = true)

$$P(J|b) = \alpha P(b) \sum_{e} P(e) \sum_{a} P(a|b,e) P(J|a) \sum_{m} P(m|a)$$

Sum over m is identically 1; M is **irrelevant** to the query



Thm 1: Y is irrelevant unless $Y \in Ancestors(\{X\} \cup \mathbf{E})$

Here,
$$X = JohnCalls$$
, $\mathbf{E} = \{Burglary\}$, and $Ancestors(\{X\} \cup \mathbf{E}) = \{Alarm, Earthquake\}$ so M is irrelevant

(Compare this to backward chaining from the query in Horn clause KBs)

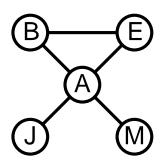
Irrelevant variables contd.

Defn: moral graph of Bayes net: marry all parents and drop arrows

Defn: A is m-separated from B by C iff separated by C in the moral graph

Thm 2: Y is irrelevant if m-separated from X by \mathbf{E}

For P(JohnCalls|Alarm=true), both Burglary and Earthquake are irrelevant



Complexity of exact inference

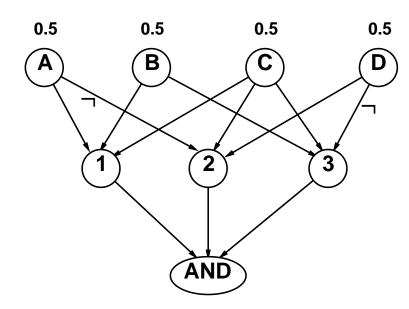
Singly connected networks (or polytrees):

- any two nodes are connected by at most one (undirected) path
- time and space cost of variable elimination are $O(d^k n)$

Multiply connected networks:

- can reduce 3SAT to exact inference \Rightarrow NP-hard
- equivalent to **counting** 3SAT models \Rightarrow #P-complete





Inference by stochastic simulation

Basic idea:

- 1) Draw N samples from a sampling distribution S
- 2) Compute an approximate posterior probability \hat{P}
- 3) Show this converges to the true probability P

Outline:

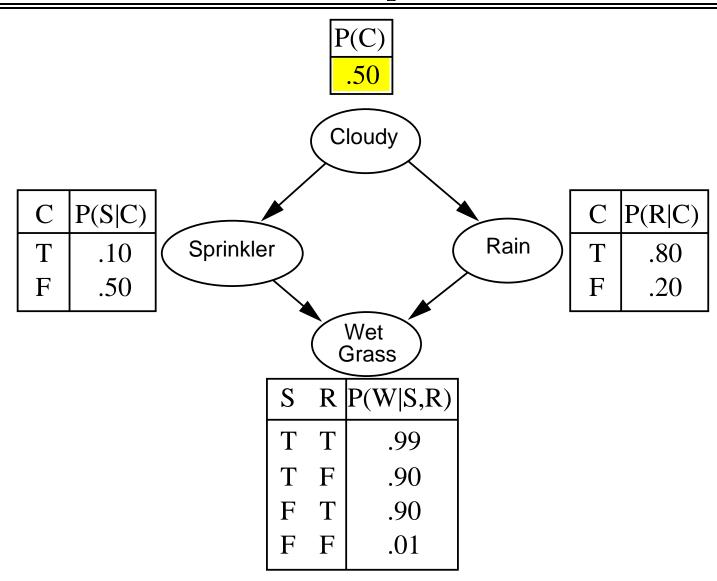
- Sampling from an empty network
- Rejection sampling: reject samples disagreeing with evidence
- Likelihood weighting: use evidence to weight samples
- Markov chain Monte Carlo (MCMC): sample from a stochastic process whose stationary distribution is the true posterior



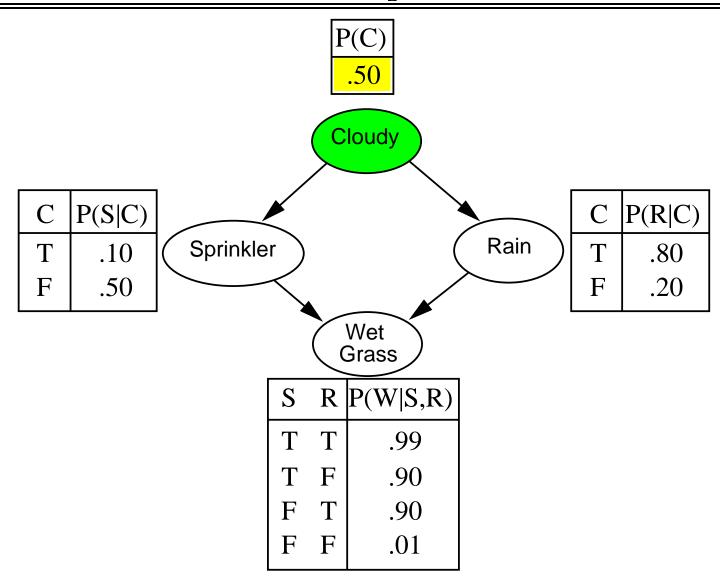
Sampling from an empty network

```
function PRIOR-SAMPLE(bn) returns an event sampled from bn inputs: bn, a belief network specifying joint distribution \mathbf{P}(X_1,\ldots,X_n) \mathbf{x}\leftarrow an event with n elements for i=1 to n do x_i\leftarrow a random sample from \mathbf{P}(X_i\mid Parents(X_i)) return \mathbf{x}
```

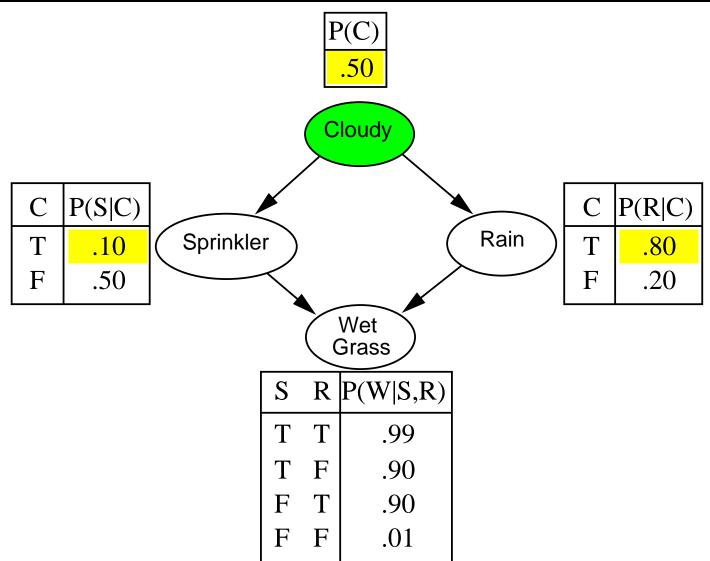
Example



Example



Example



Example P(C).50 Cloudy P(S|C)P(R|C)Rain Sprinkler .10 .80 T .50 .20 F F Wet Grass $R \mid P(W|S,R)$ S T .99 F .90

.90

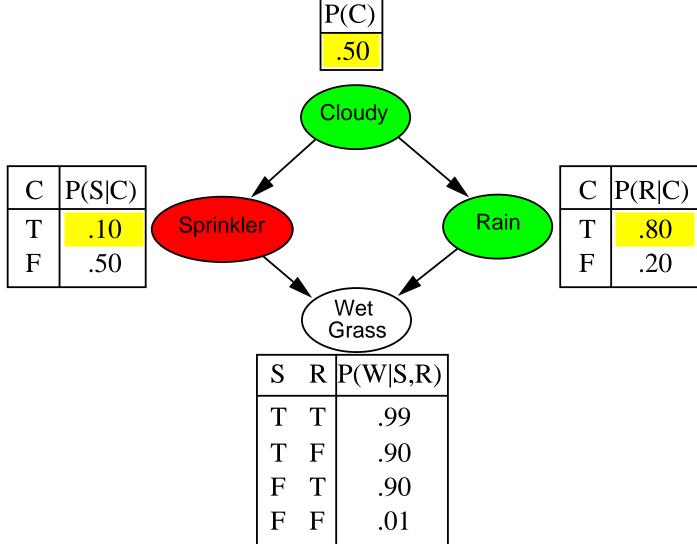
.01

F

T

F

Example P(C)



Example P(C).50 Cloudy P(S|C)P(R|C)Rain Sprinkler .10 .80 T .50 .20 F F Wet Grass $R \mid P(W|S,R)$ S T .99

F

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Example P(C).50 Cloudy P(S|C)P(R|C)Rain Sprinkler .10 .80 T .50 .20 F F Wet Grass $R \mid P(W|S,R)$ S T .99 F .90 .90 F T F .01

Sampling from an empty network contd.

Probability that PRIORSAMPLE generates a particular event

$$S_{PS}(x_1 \dots x_n) = \prod_{i=1}^n P(x_i | Parents(X_i)) = P(x_1 \dots x_n)$$

i.e., the true prior probability

E.g.,
$$S_{PS}(t, f, t, t) = 0.5 \times 0.9 \times 0.8 \times 0.9 = 0.324 = P(t, f, t, t)$$

Let $N_{PS}(x_1 \dots x_n)$ be the number of samples generated for event x_1, \dots, x_n

Then we have

$$\lim_{N \to \infty} \hat{P}(x_1, \dots, x_n) = \lim_{N \to \infty} N_{PS}(x_1, \dots, x_n) / N$$

$$= S_{PS}(x_1, \dots, x_n)$$

$$= P(x_1 \dots x_n)$$

That is, estimates derived from PRIORSAMPLE are consistent

Shorthand: $\hat{P}(x_1, \ldots, x_n) \approx P(x_1 \ldots x_n)$

Rejection sampling

 $\hat{\mathbf{P}}(X|\mathbf{e})$ estimated from samples agreeing with \mathbf{e}

```
function Rejection-Sampling(X, e, bn, N) returns an estimate of P(X|e) local variables: \mathbf{N}, a vector of counts over X, initially zero for j=1 to N do \mathbf{x}\leftarrow \text{Prior-Sample}(bn) if \mathbf{x} is consistent with \mathbf{e} then \mathbf{N}[x]\leftarrow \mathbf{N}[x]+1 where x is the value of X in \mathbf{x} return \text{Normalize}(\mathbf{N}[X])
```

E.g., estimate $\mathbf{P}(Rain|Sprinkler=true)$ using 100 samples 27 samples have Sprinkler=true Of these, 8 have Rain=true and 19 have Rain=false.

$$\hat{\mathbf{P}}(Rain|Sprinkler = true) = \text{Normalize}(\langle 8, 19 \rangle) = \langle 0.296, 0.704 \rangle$$

Similar to a basic real-world empirical estimation procedure

Analysis of rejection sampling

```
\hat{\mathbf{P}}(X|\mathbf{e}) = \alpha \mathbf{N}_{PS}(X,\mathbf{e}) (algorithm defn.)

= \mathbf{N}_{PS}(X,\mathbf{e})/N_{PS}(\mathbf{e}) (normalized by N_{PS}(\mathbf{e}))

\approx \mathbf{P}(X,\mathbf{e})/P(\mathbf{e}) (property of PRIORSAMPLE)

= \mathbf{P}(X|\mathbf{e}) (defn. of conditional probability)
```

Hence rejection sampling returns consistent posterior estimates

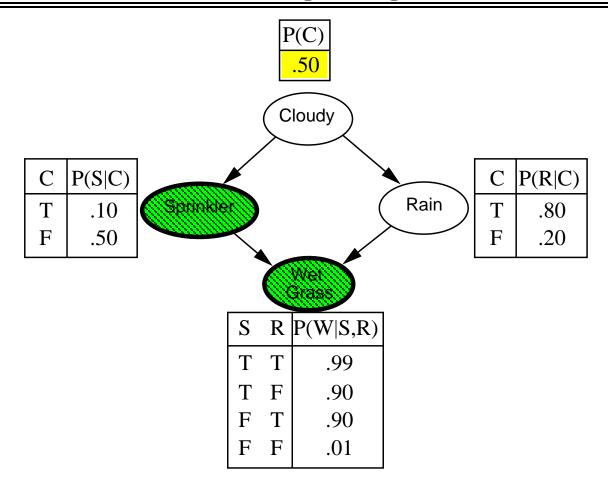
Problem: hopelessly expensive if $P(\mathbf{e})$ is small

 $P(\mathbf{e})$ drops off exponentially with number of evidence variables!

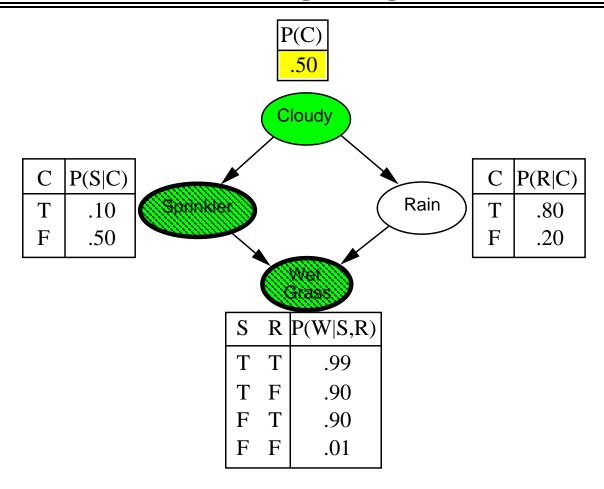
Likelihood weighting

Idea: fix evidence variables, sample only nonevidence variables, and weight each sample by the likelihood it accords the evidence

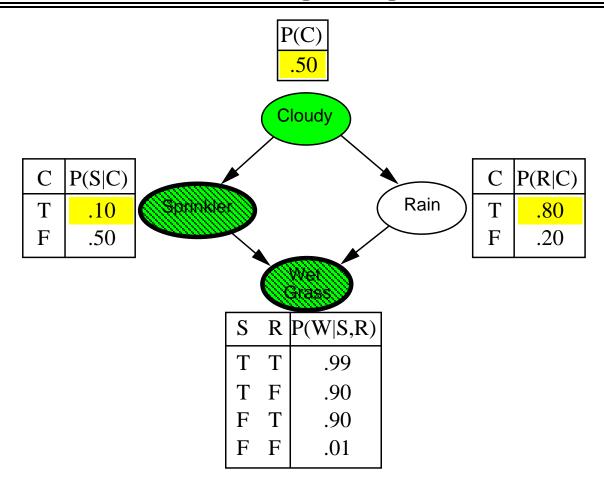
```
function LIKELIHOOD-WEIGHTING(X, \mathbf{e}, bn, N) returns an estimate of P(X|\mathbf{e})
   local variables: W, a vector of weighted counts over X, initially zero
   for j = 1 to N do
         \mathbf{x}, w \leftarrow \text{Weighted-Sample}(bn)
         \mathbf{W}[x] \leftarrow \mathbf{W}[x] + w where x is the value of X in \mathbf{x}
   return Normalize(\mathbf{W}[X])
function WEIGHTED-SAMPLE(bn, e) returns an event and a weight
   \mathbf{x} \leftarrow an event with n elements; w \leftarrow 1
   for i = 1 to n do
         if X_i has a value x_i in e
              then w \leftarrow w \times P(X_i = x_i \mid Parents(X_i))
              else x_i \leftarrow a random sample from P(X_i \mid Parents(X_i))
   return x, w
```



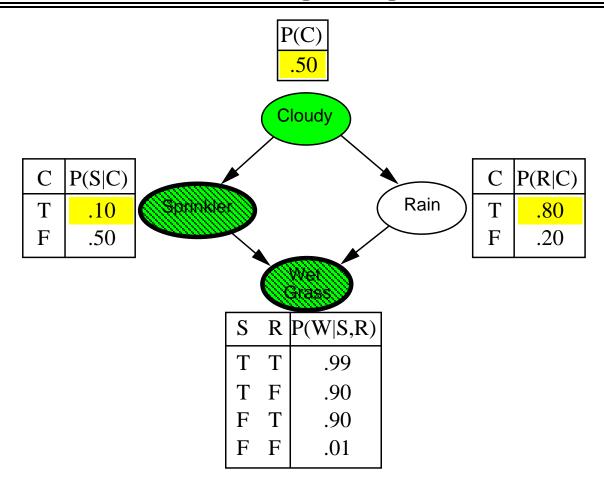
w = 1.0



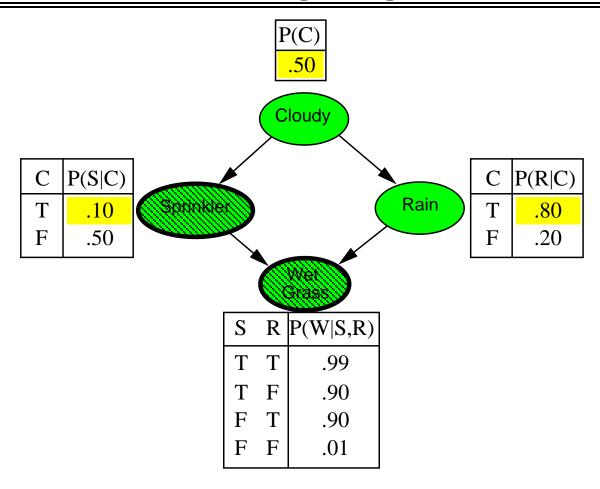
w = 1.0



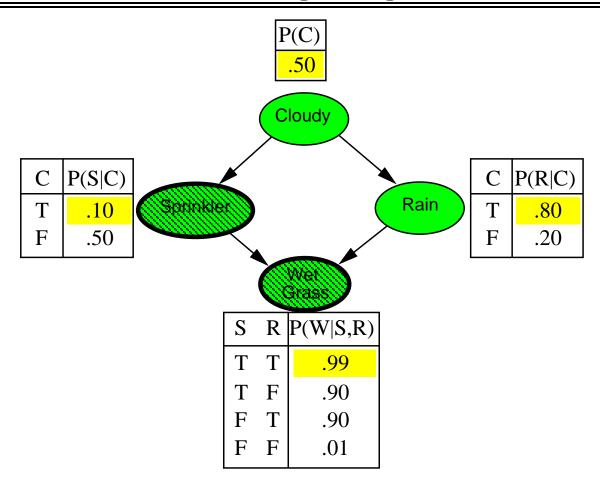
w = 1.0



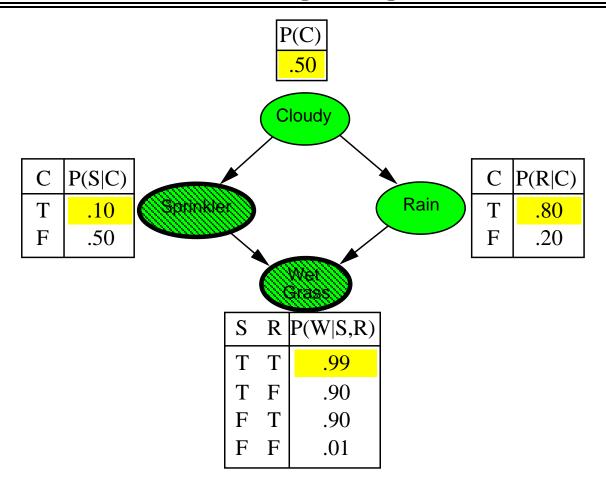
 $w = 1.0 \times 0.1$



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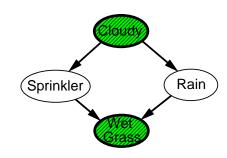
$$w = 1.0 \times 0.1 \times 0.99 = 0.099$$

Likelihood weighting analysis

Sampling probability for WEIGHTEDSAMPLE is

$$S_{WS}(\mathbf{z}, \mathbf{e}) = \prod_{i=1}^{l} P(z_i | Parents(Z_i))$$

Note: pays attention to evidence in **ancestors** only ⇒ somewhere "in between" prior and posterior distribution



Weight for a given sample z, e is

$$w(\mathbf{z}, \mathbf{e}) = \prod_{i=1}^{m} P(e_i | Parents(E_i))$$

Weighted sampling probability is

$$S_{WS}(\mathbf{z}, \mathbf{e})w(\mathbf{z}, \mathbf{e})$$

$$= \prod_{i=1}^{l} P(z_i|Parents(Z_i)) \quad \prod_{i=1}^{m} P(e_i|Parents(E_i))$$

$$= P(\mathbf{z}, \mathbf{e}) \text{ (by standard global semantics of network)}$$

Hence likelihood weighting returns consistent estimates but performance still degrades with many evidence variables because a few samples have nearly all the total weight

Approximate inference using MCMC

"State" of network = current assignment to all variables.

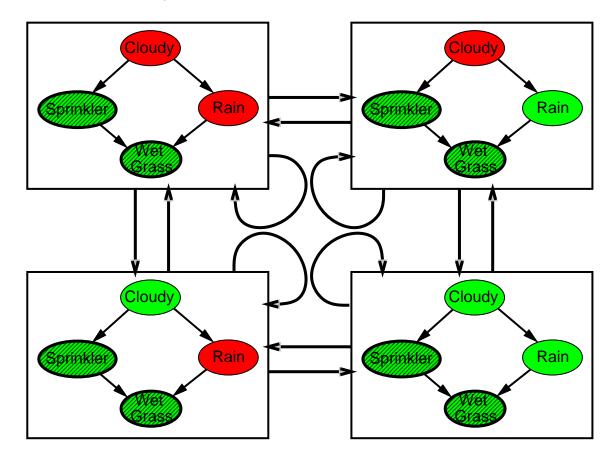
Generate next state by sampling one variable given Markov blanket Sample each variable in turn, keeping evidence fixed

```
function MCMC-Ask(X, e, bn, N) returns an estimate of P(X|e) local variables: \mathbf{N}[X], a vector of counts over X, initially zero \mathbf{Z}, the nonevidence variables in bn \mathbf{x}, the current state of the network, initially copied from e initialize \mathbf{x} with random values for the variables in \mathbf{Y} for j=1 to N do \mathbf{N}[x]\leftarrow\mathbf{N}[x]+1 where x is the value of X in \mathbf{x} for each Z_i in \mathbf{Z} do sample the value of Z_i in \mathbf{x} from \mathbf{P}(Z_i|MB(Z_i)) given the values of MB(Z_i) in \mathbf{x} return \mathrm{NORMALIZE}(\mathbf{N}[X])
```

Can also choose a variable to sample at random each time

The Markov chain

With Sprinkler = true, WetGrass = true, there are four states:



Wander about for a while, average what you see

MCMC example contd.

Estimate P(Rain|Sprinkler = true, WetGrass = true)

Sample Cloudy or Rain given its Markov blanket, repeat. Count number of times Rain is true and false in the samples.

E.g., visit 100 states 31 have Rain = true, 69 have Rain = false

$$\hat{\mathbf{P}}(Rain|Sprinkler = true, WetGrass = true) \\ = \text{Normalize}(\langle 31, 69 \rangle) = \langle 0.31, 0.69 \rangle$$

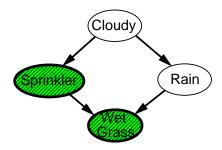
Theorem: chain approaches stationary distribution: long-run fraction of time spent in each state is exactly proportional to its posterior probability

Markov blanket sampling

Markov blanket of Cloudy is Sprinkler and Rain

Markov blanket of Rain is

Cloudy, Sprinkler, and WetGrass



Probability given the Markov blanket is calculated as follows:

$$P(x_i'|MB(X_i)) = P(x_i'|Parents(X_i)) \prod_{Z_j \in Children(X_i)} P(z_j|Parents(Z_j))$$

Easily implemented in message-passing parallel systems, brains

Main computational problems:

- 1) Difficult to tell if convergence has been achieved
- 2) Can be wasteful if Markov blanket is large:

 $P(X_i|MB(X_i))$ won't change much (law of large numbers)

Summary

Exact inference by variable elimination:

- polytime on polytrees, NP-hard on general graphs
- space = time, very sensitive to topology

Approximate inference by LW, MCMC:

- LW does poorly when there is lots of (downstream) evidence
- LW, MCMC generally insensitive to topology
- Convergence can be very slow with probabilities close to 1 or 0
- Can handle arbitrary combinations of discrete and continuous variables